## Item response theory analyses of the Reading the Mind in the Eyes Test

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## Introduction

The Reading the Mind in the Eyes test (RMET; Baron-Cohen, Wheelwright, Hill, Raste, \& Plumb, 2001) is used as a measure of theory of mind (Black \& Barnes, 2015a, 2015b; Kidd \& Castano, 2013; Panero et al., 2016), mentalizing (Kidd \& Castano, 2016), empathy (Djikic, Oatley, \& Moldoveanu, 2013), and/or perspective-taking (Bischoff and Peskin,2014), both in experimental and correlational designs, most recently related to the effects of fiction and nonfiction narratives. Some researchers categorize the items by valence (e.g., Chamorro-Premuzic \& Ali, 2010) but most treat it as a unidimensional construct. To the best of our knowledge, only one study has been done to investigate the item properties: Preti, Vellante, and Petretto (2017) applied Item Response Theory (IRT) to an Italian sample ( $\mathrm{N}=200$ ) and report evidence for a unidimensional structure, with a Rasch model as best fit The purpose of this research was to utilize IRT to analyze RMET item properties in an English speaking sample in order to (a) determine how many, if any, items could be dropped for a shorter instrument for use in future research, and (b) to estimate item parameters for use in rescoring and testing outcomes from past experiments.

## Method

## Participants

Three labs contributed a total of 365 cases to this project

- All data were from control groups (no manipulations)
- $\mathrm{N}=138$ were OU undergraduates who completed an online survey in exchange for credit.
- $\mathrm{N}=31$ were undergraduates recruited from the departmental pool at an East Coast university $\mathrm{N}=190$ participated via Amazon.com's Mechanical Turk
Data Analyses
IRTPRO and RStudio (ltm package) were used to estimate item parameters. 2PL analyses were run in both; parameters matched. R was used to test a Rasch model and obtain AIC/SBC statistics for model comparison. The full 3PL model did not converge without constraining guessing parameters. The 26 - and 22 -item tests converged without imposing contraints. Items were dropped successively based on item information curves, discrimination parameters, and fit


Item 36 (most correctan answers, mosis sensitive)



Hem 23 (most difficult)


## Results

Table
Item statistics from classic reliability analysis and two and three parameter IRT analyses.

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- 26 item version: 2 item fit statistics were significant, $r_{\alpha}=.80$
- 22 item version: no fit problems, $r_{\alpha}=.76$
- 2PL a significantly better fit than Rasch for all models, ps < . 001
- No difference between 2PL and 3PL models (full or reduced test), ps >.750; 2PL preferred for parsimony



## Discussion

Many items were very poor; no discrimination between people of similar ability. Failure to find anything greater than a small effect of manipulation in past research may reflect poor measurement.

The RMET is an easy test; only two items had difficulty statistics above the mean.

- Sample of neurotypical adults... people are good at being social primates.
- Need for more challenging measure if we want to discriminate between high levels of theory of mind ability.

Limitations \& future research
The 36 -item test is probably not unidimensional. The single factor model was not a good fit for the full test; possibly I just eliminated items that should have been part of a second factor. A simple LRT in R suggested that a two-factor model would fit the data better ( $p<.001$ ). Future research should include exploratory factor analyses.

We still don't know if this is primarily a vocabulary test.

Need to test for differential item functioning across different groups (e.g., gender, sample source).

| Model | items | $\alpha$ | $\mathrm{M}_{2}$ | $d f$ | $p$ | RMSEA | $\begin{gathered} \hline \text { Log } \\ \text { likelihood } \end{gathered}$ | aic | BC |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Rasch | 36 | . 803 |  |  |  |  | -7050 | 14174 | 14318 |
| 2 PL |  |  | 822 | 594 | <. 001 | . 03 | -6988 | 14120 | 14400 |
| 3 PL |  |  | 741 | 558 | - . 201 | . 03 | -6996 | 14208 | 14629 |
| 2 PL | 26 | . 800 | 317 | 299 | . 229 | . 01 | -4871 | 9846 | 10048 |
| 3 PL |  |  | 271 | 273 | . 516 | . 01 | -4878 | 9912 | 10216 |
| 2 PL | 22 | . 758 | 243 | 209 | . 055 | . 02 | -4170 | 8428 | 8600 |
| 3PL |  |  | 205 | 187 | . 170 | . 02 | -4175 | 8482 | 8740 |

Model fit statistics for single-factor solution, full and reduced RMET
 $\begin{array}{llllll}0.384 & .056[051,060] & .015 & 0.430 & 1480 & 1897\end{array}$ $\begin{array}{lllllllllll}26 \text { item } & 78 & 376 & 299 & 1.03 & 0.045 & .0277 .017,035] & 1.000 & 0.918 & 532 & 836\end{array}$ $\begin{array}{lllllllllll}22 \text { item } & 66 & 285 & 209 & 0.78 & 0.046 & .032[.022, .040] & .999 & 0.876 & 417 & 674\end{array}$


26 items: $4,5,7,8,9,10,11,12,13,14,15,16,18,20,22$,
$23,24,26,27,28,30,32,33,34,35,36$
22 items: $4,5,7,8,9,10,11,15,16,18,20,23,24,26,27$, $28,30,32,33,34,35,36$

